Mathematics 228 PS4 Solutions
§7.1 38
In the first case

|  | A | a |
| :---: | :---: | :---: |
| A | 0.28 | 0.42 |
| a | 0.12 | 0.18 |

$\operatorname{Pr}(A a)=0.54$
in the second case

|  | $A$ | $a$ |
| :---: | :---: | :---: |
| $A$ | 0.35 | 0.15 |
| $a$ | 0.15 | 0.35 |

§7.2 28
After elimination the $0,1,2$ probabilities are $\frac{4}{9}, \frac{3}{9}$, and $\frac{2}{9}$ respectively. We have the following joint distribution ( $\mathrm{S}=1$ means stained)

|  | 1 | 0 |
| :---: | :---: | :---: |
| 0 | $\frac{19}{45}$ | $\frac{1}{45}$ |
| 1 | $\frac{3}{10}$ | $\frac{1}{30}$ |
| 2 | $\frac{8}{45}$ | $\frac{2}{45}$ |

$E(A)=\frac{7}{9}$
$E(S)=\frac{9}{10}$
$\mathrm{E}(\mathrm{AS})=\frac{59}{90}$
So, $\operatorname{Cov}(A, S)=-\frac{2}{45}$. The negative makes sense, as it's harder to stain as the age increases, so as $A$ increases, $S$ tends to decrease.
§7.2 32
The joint distribution for $\mathrm{S}, \mathrm{P}$ is given by

|  | 8 | 9 | 5 |
| :--- | :--- | :--- | :--- |
| 10 | .7 | 0 | 0 |
| 14 | 0 | .2 | 0 |
| 18 | 0 | 0 | .1 |

$E(S)=11.6$
$\mathrm{E}(\mathrm{P})=7.9$
$E(S P)=90.2$ so $\operatorname{Cov}(S, P)=-1.44$
$E\left(S^{2}\right)=141.6 E\left(P^{2}\right)=63.5$, so $\operatorname{Var}(S)=7.04$ and $\operatorname{Var}(P)=1.09$
therefore $\rho(S, P)=-0.5198319$
This is not -1 because the bird also eats, unlike in the book example.
§7.3 36 (extra credit)
The expected revenue of the first is $\$ 1000(.6+.2)=\$ 800$
The second is $\$ 1000(.2+.4)=\$ 600$
The third is $\$ 500(.3+1.2)=\$ 750$
The total is therefore $\$ 2150$.
The variance for the first is $(\$ 1000)^{\wedge} 2(0.36)=\$ 360,000$
The second is $(\$ 1000)^{\wedge} 2(0.64)=\$ 640,000$
The third is $(\$ 500)^{\wedge} 2(0.45)=\$ 112,500$
If independent the total variance is $\$ 1,112,500$
The total standard deviation is then $\$ 3335$, which is more than the total mean. This is not very predictable business.
§7.4 52

We have $\binom{5}{3}=10$ ways to distribute medication. One is correct, so the probability of being correct is 0.1 .

This follows binomial - both alleles are independent and identical, so the probabilities are $1(.8)^{2}(.2)^{0}=0.64 \mathrm{AA}, 2(.8)^{1}(.2)^{1}=0.32 \mathrm{Aa}$, and $1(.8)^{0}(.2)^{2}=0.04$.
§7.5 (37, 39, 40) as one question
We expect $40(.2)^{\mathrm{t}}$ remaining t minutes later
The variance is $40(.2)^{\mathrm{t}}\left(1-(.2)^{\mathrm{t}}\right)$.
We take the derivative, notice that the maximum does not occur at 0 , and find where the derivative is zero to be the maximum, $t=.8126795911$ minutes, but this is discrete, not continuous, so we check and find that 1 is larger than 0 , so the maximum is at one minute.

The coefficient of variation is $\frac{\sqrt{40\left(.2^{t}-.04^{t}\right)}}{40(.2)^{t}}$ a graph reveals that this function is increasing.
§7.6 (59-64) as one question
Following the text on pp. 652-3, we have
$\frac{d P}{d t}=-t P$
separate
$\int \frac{d P}{P}=\int-t d t$
$\ln (P)=-\frac{1}{2} \mathrm{t}^{2}+\mathrm{C}$
$P=e^{-\frac{1}{2} t^{2}}$ is the probability surviving to time $t$. Graph it on $[0,4] \times[0,1]$.
The cdf is $1-P=1-e^{-\frac{1}{2} t^{2}}$ Differentiate to find the pdf $t e^{-\frac{1}{2} t^{2}}$
$P(1)=.6065306597$ is the probability surviving to age 1
$P(2) / P(1)=.2231301601$ is the probability surviving to age 2
conditional on surviving to age 1 . This is lower than the previous answer because death rate increases with age. This would not be the case for a constant death rate.
$\sum_{\mathrm{k}=0}^{\infty} \frac{1}{\mathrm{k!}}(1-\mathrm{q})^{\mathrm{k}} \mathrm{e}^{-\lambda \mathrm{t}}(\lambda \mathrm{t})^{\mathrm{k}}$
k corresponds to how many times the organism has been attacked.
Notice the $(1-q)^{\mathrm{k}}$ corresponds to attacks failing. Notice also that this does work for $\mathrm{k}=0$ as desired.
$\S 7.836$
The expected value of the annual change is -3.5 . The variance is 28.45 . So, the expected value of the 10 year average is also -3.5 , but the variance is 2.845 . So, the pdf is $\frac{1}{\sqrt{2 \pi(2.845)}} \mathrm{e}^{-(x+3.5)^{2} / 2(2.845)}$. The maximum average possible is 2 , the minimum possible is -10 . This should gives us a clue for where to plot. $[-2,10] \times[0,0.3]$ looks good. Not much above 0, maybe $5 \%$ ? How do we feel about the approximation? Looks pretty ok, but could argue that it isn't.

## $\S 7.950$

Using the binomial distribution the mean is 4 and the variance is 3.992 . We may then use the normal approximation. The appropriate pdf is $\frac{1}{\sqrt{2 \pi 3.992}} \mathrm{e}^{-(\mathrm{x}-4)^{2} / 2(3.992)}$, which we integrate from .5 to 1.5 to get 0.065513

If you use the Poisson approximation to the binomial (as Adler may suggest) the variance is 4 instead of 3.992 yielding 0.0655906

## §8.1 44

We can do this either using Poisson (as the book does, and is hinted by the use of $\lambda$ ) or using binomial. I'll do Poisson first:

The probability of 12 out of 200 nonsynonymous is given by $\frac{\mathrm{e}^{-\lambda 200}(200 \lambda)^{12}}{12!}$ and the probability of 15 out of 100 synonymous is given by $\frac{e^{-3 \lambda 100}(100 \cdot 3 \lambda)^{15}}{15!}$. So the probability both happen is $L(\lambda)=\left(\frac{\mathrm{e}^{-\lambda 200}(200 \lambda)^{12}}{12!}\right)\left(\frac{e^{-3 \lambda 100}(100 \cdot 3 \lambda)^{15}}{15!}\right)$
$=\frac{e^{-500 \lambda} 200^{12} 300^{15} \lambda^{27}}{15!12!}$. The constants don't affect our answer, so we write them hereafter
$L^{\prime}(\lambda)=-500 e^{-500 \lambda} \lambda^{27}+27 e^{-500 \lambda} \lambda^{26}=\lambda^{26} e^{-500 \lambda}(27-500 \lambda)=0$
$L(0)=0, L\left(\frac{27}{500}\right)=0.000000116622 \neq 0$, so $\frac{27}{500}=0.054$ is the mle.
Seems to fit the data.
Now redoing with binomial. The probability of 12 out of 200 nonsynonymous is given by $\binom{200}{12} \lambda^{12}(1-\lambda)^{188}$ and the probability of 15 out of 100 synonymous is given by $\binom{100}{15}(3 \lambda)^{15}(1-3 \lambda)^{85}$. So the probability both happen is $L(\lambda)=\binom{200}{12} \lambda^{12}(1-\lambda)^{188}\binom{100}{15}(3 \lambda)^{15}(1-3 \lambda)^{85}$

$$
=\binom{200}{12}\binom{100}{15} \lambda^{27}(1-\lambda)^{188} 3^{15}(1-3 \lambda)^{85} \text {. The constants don't affect our }
$$

answer, so we don't write them hereafter. Use the product rule twice to

$$
L^{\prime}(\lambda)=-\lambda^{26}(1-\lambda)^{187}(1-3 \lambda)^{84}\left(900 \lambda^{2}-551 \lambda+27\right)=0
$$

$\mathrm{L}(1)=\mathrm{L}(0)=\mathrm{L}(1 / 3)=0,0.55850764$ is unrealistic, as it produces $\lambda_{\mathrm{s}}<0$ Since $\mathrm{L}(0.053714573208)=0.0116009>0,0.053714573208$ is the mle. Close to the previous, a little lower.
§8.3 (33-36) as one question
For each single measurement, we have $p(1-p)$ as a variance, in this case the variance is $\frac{1}{4}$.

| sample | probability | $\overline{\mathrm{x}}$ | msd | sample variance |
| :---: | :---: | :---: | :---: | :---: |
| $(0,0)$ | $\frac{1}{4}$ | 0 | 0 | 0 |
| $(1,0)$ | $\frac{1}{2}$ | $\frac{1}{2}$ | $\frac{1}{4}$ | $\frac{1}{2}$ |
| $(1,1)$ | $\frac{1}{4}$ | 0 | 0 | 0 |

See that the expected value of $\mathrm{msd}=\frac{1}{8}$, whereas the expected value of the sample variance is $\frac{1}{4}$, which is the actual variance. This is an example of why we use the sample variance with $n-1$ instead of $n$.

## §8.4 30

The survival time is exponentially distributed with a mean of 100 h . Because the mean of an exponential distribution is $\frac{1}{\lambda}$, we have an exponential distribution with parameter 0.01 . We seek a survival time, s, such that the probability of it occurring is less than or equal to 0.05 , i.e. such that $\int_{0}^{s} 0.01 \mathrm{e}^{-0.01 t} \mathrm{dt} \leq 0.05$

This yields $1-\mathrm{e}^{-0.01 \mathrm{~s}} \leq 0.05$
solving for s produces $\mathrm{s} \leq 5.129 \mathrm{~h}$. It must be pretty small to be significantly different from 100 . It's worth recalling that the standard deviation of the exponential distribution with mean 100 is 100 , so it's a pretty spread out distribution. Also that the mode value is 0 , so getting low values isn't entirely uncommon.

